

A NOTE ON THE COMPOSITION OF DISTRIBUTIONS

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Let F be a distribution and let f be a locally summable function. The distribution $F(f)$ is defined as the neutrix limit of the sequence $\{F_n(f)\}$, where $F_n(x) = F(x) * \delta_n(x)$ and $\{\delta_n(x)\}$ is a certain sequence of infinitely differentiable functions converging to the Dirac delta-function $\delta(x)$. The distribution $(x_+)^{2r-1/2}$ is evaluated for $r = 1, 2, \dots$

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In the following, we let \mathcal{D} be the space of all infinitely differentiable functions with compact support and let \mathcal{D}' be the space of distributions on \mathcal{D} . Further, we let ρ be an infinitely differentiable function having the properties

(i) $\rho(x) = 0, |x| \geq 1,$

(ii) $\rho(x) \geq 0,$

(iii) $\rho(x) = \rho(-x)$

and

(iv) $\int_{-1}^1 \rho(x) dx = 1.$

Putting $\delta_n(x) = n\rho(nx)$ for $n = 1, 2, \dots$, it follows that $\{\delta_n(x)\}$ is a regular sequence of infinitely differentiable functions converging to the Dirac delta-function $\delta(x)$ in the sense that

$$\lim_{n \rightarrow \infty} \langle \delta_n(x), \varphi(x) \rangle = \langle \delta(x), \varphi(x) \rangle$$

for all φ in \mathcal{D} , see Gel'fand and Shilov⁷.

If now F is an arbitrary distribution in \mathcal{D}' , we define

$$F_n(x) = (F * \delta_n)(x) = \langle F(t), \delta_n(x-t) \rangle$$

for $n = 1, 2, \dots$. It follows that $\{F_n(x)\}$ is a regular sequence of infinitely differentiable functions converging to F .

The following definition was given in [2].

Definition 1 — Let F be a distribution in \mathcal{D}' and let f be a locally summable function. We say that the distribution $F(f(x))$ exists and is equal to the distribution $h(x)$ in \mathcal{D}' on the interval (a, b) if

$$N\text{-}\lim_{n \rightarrow \infty} \int_a^b F_n(f(x)) \varphi(x) dx = \langle h(x), \varphi(x) \rangle$$

for all φ in \mathcal{D} with support contained in the interval (a, b) , where

$$F_n(x) = (F * \delta_n)(x)$$

and N is the neutrix having domain $N' = \{1, 2, \dots, n, \dots\}$ and range N'' the real numbers with negligible functions finite linear sums of the functions

$$n^\lambda \ln^{r-1} n, \ln^r n : \lambda > 0, r = 1, 2, \dots$$

and all functions which converge to zero in the normal sense as n tends to infinity, see van der Corput¹.

This definition was later extended in [6] to cover the case when f is also a distribution.

Definition 2 — Let F and f be distributions in \mathcal{D}' . We say that the distribution $D(f(x))$ exists and is equal to the distribution $h(x)$ in \mathcal{D}' on the interval (a, b) if

$$N\text{-}\lim_{n \rightarrow \infty} \left[N\text{-}\lim_{n \rightarrow \infty} \int_a^b F_n(f_m(x)) \varphi(x) dx \right] = \langle h(x), \varphi(x) \rangle$$

for all φ in \mathcal{D} with support contained in the interval (a, b) , where

$$F_n(x) = (F * \delta_n)(x), f_m(x) = (f * \delta_m)(x)$$

and N is the neutrix defined above.

The following theorems were proved in [2] and [3] respectively:

Theorem 1 — The distributions $(x_{-}^{\mu})_{-}^{\lambda}$ and $(x_{+}^{\mu})_{-}^{\lambda}$ exists and

$$(x_{-}^{\mu})_{-}^{\lambda} = (x_{+}^{\mu})_{-}^{\lambda} = 0$$

for $\mu > 0$, $\lambda \neq -1, 2, \dots$ and $\lambda\mu \neq -1, -2, \dots$ and

$$(x_{-}^{\mu})_{-}^{\lambda} = (-1)^{\lambda\mu} (x_{+}^{\mu})_{-}^{\lambda} = \frac{\pi \operatorname{cosec}(\pi\lambda)}{2\mu(-\lambda\mu-1)!} \delta^{(-\lambda\mu-1)}(x)$$

for $\mu > 0$, $\lambda \neq -1, -2, \dots$ and $\lambda\mu = -1, -2, \dots$

Theorem 2 — The distribution $(x^2)_{+}^{-s-1/2}$ exists and

$$(x^2)_{+}^{-s-1/2} = |x|^{-2s-1} + \frac{2}{(2s)!} [\ln 2 - c(\rho)] \delta^{(2s)}(x),$$

for $s = 0, 1, 2, \dots$

Further results were later proved in [4], [5] and [6].

Before proving our main theorem we note the following lemma which can be proved easily.

Lemma 1 — Let φ be a function in \mathcal{D} with support contained in the interval $[-1, 1]$. Then

$$\begin{aligned} \langle x_{+}^{-r}, \varphi(x) \rangle &= \int_0^1 x^{-r} \left[\varphi(x) - \sum_{i=0}^{r-1} \frac{\varphi^{(i)}(0)}{i!} x^i \right] dx - \sum_{i=0}^{r-2} \frac{\varphi^{(i)}(0)}{i!(r-i-1)} \\ &\quad - \frac{\varphi^{(r-1)}(0)}{(r-1)!} \varphi^{(r-1)}(0), \end{aligned}$$

for $r = 1, 2, \dots$, where

$$\varphi(r) = \begin{cases} \sum_{i=1}^r 1/i, & r \geq 1, \\ 0, & r = 0. \end{cases}$$

In the above lemma, the distribution x_{+}^{-r} is defined by

$$x_{+}^{-r} = \frac{(-1)^{r-1}}{(r-1)!} (\ln x_{+})^{(r)},$$

for $r = 1, 2, \dots$ and not as in Gel'fand and Shilov⁷.

We now prove the following theorem.

Theorem 3 — Let $F(x)$ denote the function $x_{+}^{-1/2}$. Then the distribution $F(x_{+}^{2r})$ exists and

$$F(x_{+}^{2r}) = x_{+}^{-r} + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\varphi(r-1)]}{r!} \delta^{(r-1)}(x), \quad \dots (1)$$

for $r = 1, 2, \dots$

PROOF : We have

$$F_n(x) = \begin{cases} \int_{-1/n}^{1/n} (x-t)^{-1/2} \delta_n(t) dt, & x > 1/n, \\ x \int_{-1/n}^{1/n} (x-t)^{-1/2} \delta_n(t) dt, & -1/n < x < 1/n, \\ 0, & x < -1/n. \end{cases}$$

It follows that

$$\begin{aligned} \int_{-1}^1 F_n(x_+^{2r}) x^i dx &= \int_0^1 F_n(x^{2r}) x^i dx + \int_{-1}^0 F_n(0) x^i dx \\ &= \int_0^{n^{-1/2r}} \int_{-1/n}^{x^{2r}} (x^{2r}-t)^{-1/2} x^i \delta_n(t) dt dx \\ &\quad + \int_{n^{-1/2r}}^1 \int_{-1/n}^{1/n} (x^{2r}-t)^{-1/2} x^i \delta_n(t) dt dx \\ &\quad + \int_{-1}^0 \int_{-1/n}^0 (-t)^{-1/2} x^i \delta_n(t) dt dx \\ &= \int_0^{1/n} \delta_n(t) \int_{t^{1/2r}}^1 (x^{2r}-t)^{-1/2} x^i dx dt \\ &\quad + \int_{-1/n}^0 \delta_n(t) \int_0^1 (x^{2r}-t)^{-1/2} x^i dx dt \\ &\quad + \frac{(-1)^i}{i+1} \int_{-1/n}^0 (-t)^{-1/2} \delta_n(t) dt \quad \dots (2) \\ &= \int_0^{1/n} \delta_n(t) \int_{t^{1/2r}}^1 [(x^{2r}-t)^{-1/2} + (x^{2r}+t)^{-1/2}] x^i dx dt \\ &\quad + \int_0^{1/n} \delta_n(t) \int_0^{t^{1/2r}} (x^{2r}+t)^{-1/2} x^i dx dt \end{aligned}$$

$$\begin{aligned}
 & + \frac{(-1)^i}{i+1} \int_0^{1/n} t^{-1/2} \delta_n(t) dt \\
 & = I_1 + I_2 + I_3.
 \end{aligned}$$

Putting $nt = u$ and $n^{1/2r}x = v$, we have

$$\begin{aligned}
 I_1 &= n^{(r-i-1)/2r} \int_0^1 \rho(u) \int_{u^{1/2r}}^{n^{1/2r}} [(v^{2r}-u)^{-1/2} + (v^{2r}+u)^{-1/2}] v^i dv du \\
 &= 2n^{(r-i-1)/2r} \sum_{k=0}^{\infty} \binom{-\frac{1}{2}}{2k} \int_0^1 u^{2k} \rho(u) \int_{u^{1/2r}}^{n^{1/2r}} v^{-4rk-r+i} dv du \\
 &= 2n^{(r-i-1/2r)} \sum_{k=0}^{\infty} \binom{-\frac{1}{2}}{2k} \int_0^1 \frac{n^{(-4rk-r+i+1)/2r} - u^{(-4rk-r+i+1)/2r}}{-4rk-r+i+1} u^{2k} \rho(u) du
 \end{aligned}$$

and it follows that

$$N\text{-}\lim_{n \rightarrow \infty} I_1 = -\frac{2}{r-i-1} \int_0^1 \rho(u) du = -\frac{1}{r-i-1} \tag{3}$$

for $i = 0, 1, \dots, r - 2$.

Next we have

$$I_2 = n^{(r-i-1)/2r} \int_0^1 \rho(u) \int_0^{u^{1/2r}} (v^{2r}+u)^{-1/2} v^i dv du$$

and it follows that

$$N\text{-}\lim_{n \rightarrow \infty} I_2 = 0 \tag{4}$$

for $i = 0, 1, \dots, r - 2$.

Finally we have

$$I_3 = \frac{(-1)^i n^{1/2}}{i+1} \int_0^1 \rho(u) du$$

and it follows that

$$N\text{-}\lim_{n \rightarrow \infty} I_3 = 0 \tag{5}$$

for $i = 0, 1, \dots, r - 2$.

It now follows from eqs. (3), (4) and (5) that

$$\int_{-1}^1 F_n(x_+^{2r}) x^i dx = -\frac{1}{r-i-1} \tag{6}$$

for $i = 0, 1, \dots, r - 2$.

We now consider the case $i = r - 1$. We have

$$\int_{t^{1/2r}}^1 (x^{2r} - t)^{-1/2} x^{r-1} dx = \frac{\ln [1 + (1-t)^{1/2}] - \ln t^{1/2}}{r},$$

$$\int_0^1 (x^{2r} - t)^{-1/2} x^{r-1} dx = \frac{\ln [1 + (1-t)^{1/2}] - \ln |t|^{1/2}}{r}$$

and
$$\int_{-1/n}^0 (-t)^{-1/2} \delta_n(t) dt = n^{1/2} \int_{-1}^0 (-u)^{-1/2} \rho(u) du.$$

It follows from eq. (2) that

$$\begin{aligned} \int_{-1}^1 F_n(x_+^{2r}) x^{r-1} dx &= \frac{1}{r} \int_{-1/n}^{1/n} \ln [1 + (1-t)^{1/2}] \delta_n(t) dt \\ &\quad - \frac{1}{2} \int_{-1/n}^0 (-t)^{-1/2} \delta_n(t) dt = n^{1/2} \int_{-1}^0 (-u)^{-1/2} \rho(u) du \end{aligned}$$

and so

$$N\text{-}\lim_{n \rightarrow \infty} \int_{-1}^1 F_n(x_+^{2r}) x^{r-1} dx = \frac{\ln 2 - c(\rho)}{r}, \tag{7}$$

since

$$\int_{-1/n}^{1/n} \ln |t| \delta_n(t) dt = \int_{-1}^1 \ln |u| \rho(u) du - \ln n \int_{-1}^1 \rho(u) du.$$

When $i = r$, we have

$$\int_0^{n^{-1/2r}} |x^r F_n(x_+^{2r})| dx = \int_0^{n^{-1/2r}} \int_{-1/n}^{x^{2r}} x^r (x^{2r} - t)^{-1/2} \delta_n(t) dt dx$$

$$\begin{aligned}
 &= n^{-1/2r} \int_0^1 \int_{-1}^{v^{2r}} v^r (v^{2r} - u)^{-1/2} \rho(u) \, du \, dv \\
 &= O(n^{-1/2r})
 \end{aligned}$$

and it follows that if ψ is an arbitrary continuous function then

$$\lim_{n \rightarrow \infty} \int_0^{n^{-1/2r}} x^r F_n(x_+^{2r}) \psi(x) \, dx = 0. \tag{8}$$

Further,

$$\begin{aligned}
 \int_{-1}^0 x^r F_n(x_+^{2r}) \psi(x) \, dx &= \int_{-1}^0 x^r \psi(x) \int_{-1/n}^0 (-t)^{-1/2} \delta_n(t) \, dt \, dx \\
 &= n^{1/2} \int_{-1}^0 x^r \psi(x) \int_{-1}^0 (-u)^{-1/2} \rho(u) \, du \, dx
 \end{aligned}$$

and it follows that

$$N\text{-}\lim_{n \rightarrow \infty} \int_{-1}^0 x^r F_n(x_+^{2r}) \psi(x) \, dx = 0. \tag{9}$$

If now η is chosen so that $n^{-1/2r} < \eta < 1$, we have on making the substitution $nt = u$ again

$$\begin{aligned}
 \int_{n^{-1/2r}}^{\eta} |x^r F_n(x_+^{2r})| \, dx &= \int_{n^{-1/2r}}^{\eta} \int_{-1/n}^{1/n} x^r (x^{2r} - t)^{-1/2} \delta_n(t) \, dx \, dt \\
 &= \sum_{i=0}^{\infty} \binom{-\frac{1}{2}}{i} \int_{-1}^1 \rho(u) \int_{n^{-1/2r}}^{\eta} \frac{(-u)^i}{n^i x^{2ri}} \, dx \, du \\
 &= \sum_{i=0}^{\infty} \binom{-\frac{1}{2}}{i} \int_{-1}^1 \rho(u) \frac{(-u)^i n^{-i} (\eta^{1-2ri} - n^{-1/2r})}{1-2ri} \, du.
 \end{aligned}$$

It follows that

$$\lim_{n \rightarrow \infty} \int_{n^{-1/2r}}^{\eta} |x^r F_n(x_+^{2r})| \, dx = \eta$$

and so if ψ is a continuous function then

$$\lim_{n \rightarrow \infty} \int_{n^{-1/2r}}^{\eta} |x^r F_n(x_+^{2r}) \psi(x)| dx = O(\eta). \quad \dots (10)$$

Now let φ be an arbitrary function \mathcal{D} with support contained in the interval $[-1, 1]$. By Taylor's Theorem we have

$$\varphi(x) = \sum_{i=0}^{r-1} \frac{\varphi^{(i)}(0)}{i!} x^i + \frac{\varphi^{(r)}(\xi x)}{r!} x^r,$$

where $0 < \xi < 1$. Then

$$\begin{aligned} \langle F_n(x_+^{2r}), \varphi(x) \rangle &= \int_{-1}^1 F_n(x_+^{2r}) \varphi(x) dx \\ &= \sum_{i=0}^{r-1} \frac{\varphi^{(i)}(0)}{i!} \int_{-1}^1 F_n(x_+^{2r}) x^i dx + \frac{1}{r!} \int_{-1}^0 x^r F_n(x_+^{2r}) \varphi^{(r)}(\xi x) dx \\ &\quad + \frac{1}{r!} \int_0^{n^{-1/2r}} x^r F_n(x_+^{2r}) \varphi^{(r)}(\xi x) dx + \frac{1}{r!} \int_{n^{-1/2r}}^{\eta} x^r F_n(x_+^{2r}) \varphi^{(r)}(\xi x) dx \\ &\quad + \frac{1}{r!} \int_{\eta}^1 x^r F_n(x_+^{2r}) \varphi^{(r)}(\xi x) dx. \end{aligned}$$

Taking the neutrix limit as n tends to infinity, using eqs. (6)-(10), and noting that the sequence $\{x^r F_n(x_+^{2r})\}$ converges uniformly to 1 on the interval $[\eta, 1]$, it follows that

$$\begin{aligned} N\text{-}\lim_{n \rightarrow \infty} \langle F_n(x_+^{2r}), \varphi(x) \rangle &= - \sum_{i=0}^{r-2} \frac{\varphi^{(i)}(0)}{i!(r-i-1)} + \frac{\ln 2 - c(\rho)}{r!} \varphi^{(r-1)}(0) \\ &\quad + \int_{\eta}^1 \varphi^{(r)}(\xi x) dx + O(\eta) \\ &= - \sum_{i=0}^{r-2} \frac{\varphi^{(i)}(0)}{i!(r-i-1)} + \frac{\ln 2 - c(\rho)}{r!} \varphi^{(r-1)}(0) \\ &\quad + \int_0^1 \varphi^{(r)}(\xi x) dx, \end{aligned}$$

since η can be made arbitrarily small. Thus, on using Lemma 1, we have

$$\begin{aligned}
 N\text{-}\lim_{n \rightarrow \infty} \langle F_n(x_+^{2r}), \varphi(x) \rangle &= \int_0^1 x^{-r} \left[\varphi(x) - \sum_{i=0}^{r-1} \frac{\varphi^{(i)}(0)}{i!} x^i \right] dx - \sum_{i=0}^{r-2} \frac{\varphi^{(i)}(0)}{i!(r-i-1)} \\
 &\quad + \frac{\ln 2 - c(\rho)}{r!} \varphi^{(r-1)}(0) \\
 &= \langle x_+^{-r}, \varphi(x) \rangle + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\phi(r-1)]}{r!} \langle \delta^{(r-1)}(x), \varphi(x) \rangle,
 \end{aligned}$$

proving eq. (1) on the interval $[-1, 1]$. However, since $F(x_+^{2r}) = x_+^{-r}$ on any closed interval not containing the origin, we have proved eq. (1) on the real line.

Corollary 3.1 — The distribution $F(x_-^{2r})$ exists and

$$F(x_-^{2r}) = x_-^{-r} + \frac{\ln 2 - c(\rho) + r\phi(r-1)}{r!} \delta^{(r-1)}(x), \quad \dots (11)$$

for $r = 1, 2, \dots$

PROOF : Replacing x by $-x$ eq. (1), we have

$$F((-x)_+^{2r}) = (-x)_+^{-r} + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\phi(r-1)]}{r!} \delta^{(r-1)}(-x)$$

and so

$$F(x_-^{2r}) = x_-^{-r} + \frac{\ln 2 - c(\rho) + r\phi(r-1)}{r!} \delta^{(r-1)}(x),$$

proving eq. (11).

Corollary 3.2 — Let G denote the function $x_-^{1/2}$. Then the distribution $G(-x_+^{2r})$ exists and

$$G(-x_+^{2r}) = x_+^{-r} + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\phi(r-1)]}{r!} \delta^{(r-1)}(x), \quad \dots (12)$$

for $r = 1, 2, \dots$

PROOF : We have $G(-x) = F(x)$ and so

$$G(-x_+^{2r}) = F(x_+^{2r}) = x_+^{-r} + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\phi(r-1)]}{r!} \delta(x),$$

from equation (1), proving eq. (12).

Corollary 3.3 — The distribution $G(-x_-^{2r})$ exists and

$$G(-x_-^{2r}) = x_-^{-r} + \frac{\ln 2 - c(\rho) + r\phi(r-1)}{r!} \delta^{(r-1)}(x), \quad \dots (13)$$

for $r = 1, 2, \dots$

PROOF : Replacing x by $-x$ eq. (12) we have

$$G(-(-x)_+^{2r}) = (-x)_+^{-r} + \frac{(-1)^{r-1} [\ln 2 - c(\rho) + r\phi(r-1)]}{r!} \delta^{(r-1)}(-x)$$

and so

$$G(-x_-^{2r}) = x_-^{-r} + \frac{\ln 2 - c(\rho) + r\phi(r-1)}{r!} \delta^{(r-1)}(x),$$

proving eq. (13).

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