

STATISTICAL DECOMPOSITION OF DIRECTED DIVERGENCE DUE TO A NUMBER OF FACTORS AND INTERACTIONS AMONG THEM

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The directed divergence of one set of data from another set, arising due to a number of factors and the interactions among them has been decomposed into 'between' and 'within' various levels of each factor and each interaction, on the same lines which the corresponding decomposition for a single factor was done by Theil^{19,20} and Kapur¹⁰.

1. INTRODUCTION

One of the most important and useful properties of Shannon's¹⁷ measure of entropy is that the total entropy can be decomposed into entropy between groups and entropy within groups. This statistical decomposition principle has been most forcefully exploited by Theil²⁰ in applications of entropy and divergence measures to the study of income inequalities, racial segregation, industrial concentration, political alignment, combination of statements in balance sheets, brand preferences in marketing and so on. In his applications, Theil²⁰ also used the statistical decomposition properties of measure of directed divergence due to Kullback and Leibler¹².

Kapur¹⁰ in a recent paper has discussed the statistical decomposition of measures of entropy and directed divergence due to Shannon¹⁷, Havrda and Charvat⁸, Renyi¹¹, Sharma and Teneja¹⁸, Ferrari² and Kapur⁵⁻⁸. Kapur¹⁰ has also extended this principle of decomposition to measures of inaccuracy due to Kerridge¹¹ and Kapur⁹ and to measures of useful information due to Belis and Guiasu¹. He has also shown that the decomposition of the total SS in the analysis of variance can be interpreted as analysis of directed divergence of order two.

In an alternative approach, Rao^{13,15} has developed techniques for decomposition of diversity (DEDIV), apportionment of diversity (APDIV) and analysis of diversity (ANYDIV) motivated by problems of anthropology, genetics, economics, sociology and ecology.

In the present paper, we extend Kapur's¹⁰ discussion to the case of a number of factors and of interactions among them.

2. STATISTICAL DECOMPOSITION DUE TO TWO-FACTORS

Let x_{ij} and y_{ij} be the values in the two sets of data corresponding to the i th level of the first factor and j th level of the second factor ($i = 1, 2, \dots, m; j = 1, 2, \dots, n$). Let

$$\begin{aligned} \sum_{j=1}^n x_{ij} &= T_{i.}, \quad \sum_{i=1}^m x_{ij} = T_{.j}, \quad \sum_{i=1}^m \sum_{j=1}^n x_{ij} = \sum_{i=1}^m T_{i.} \\ &= \sum_{j=1}^n T_{.j} = T_{..} \quad \dots(1) \end{aligned}$$

$$\sum_{j=1}^n y_{ij} = U_{i.}, \quad \sum_{i=1}^m y_{ij} = U_{.j}, \quad \sum_{i=1}^m \sum_{j=1}^n y_{ij} = \sum_{i=1}^m U_{i.} = \sum_{j=1}^n U_{.j} = U_{..} \quad \dots(2)$$

then

$$\begin{aligned} D &= \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{T_{..}} \ln \frac{x_{ij}/T_{..}}{y_{ij}/U_{..}} = \frac{1}{2} \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{T_{..}} \\ &\times \ln \frac{x_{ij}/T_{i.} \cdot x_{ij}/T_{.j} \cdot T_{i.}/T_{..} \cdot T_{.j}/T_{..}}{y_{ij}/U_{i.} \cdot y_{ij}/U_{.j} \cdot U_{i.}/U_{..} \cdot U_{.j}/U_{..}} \\ &= \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{x_{ij}/T_{i.}}{y_{ij}/U_{i.}} \\ &+ \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \sum_{i=1}^m \frac{x_{ij}}{T_{.j}} \ln \frac{x_{ij}/T_{.j}}{y_{ij}/U_{.j}} \\ &+ \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{T_{i.}/T_{..}}{U_{i.}/U_{..}} \\ &+ \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \sum_{i=1}^m \frac{x_{ij}}{T_{.j}} \ln \frac{T_{.j}/T_{..}}{U_{.j}/U_{..}} = \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} D_{i.} \\ &+ \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} D_{.j} + \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \ln \frac{T_{i.}/T_{..}}{U_{i.}/U_{..}} \end{aligned}$$

(equation continued on p. 1409)

$$\begin{aligned}
& + \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \ln \frac{T_{.j}/T_{..}}{U_{.j}/U_{..}} = \frac{1}{2} D_{w_1} + \frac{1}{2} D_{w_2} \\
& + \frac{1}{2} D_{B_1} + \frac{1}{2} D_{B_2} \quad \dots(3)
\end{aligned}$$

where

$$D_{i.} = \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{x_{ij}/T_{i.}}{y_{ij}/U_{i.}}, D_{.j} = \sum_{i=1}^m \frac{x_{ij}}{T_{.j}} \ln \frac{x_{ij}/T_{.j}}{y_{ij}/U_{.j}} \quad \dots(4)$$

$$D_{w_1} = \sum_{i=1}^m \frac{T_{i.}}{T_{..}} D_{i.}, D_{w_2} = \sum_{j=1}^n \frac{T_{.j}}{T_{..}} D_{.j} \quad \dots(5)$$

$$D_{B_1} = \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \ln \frac{T_{i.}/T_{..}}{T_{i.}/U_{..}}, D_{B_2} = \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \ln \frac{T_{.j}/T_{..}}{T_{.j}/U_{..}} \quad \dots(6)$$

so that $D_{i.}$, $D_{.j}$ are the directed divergences within the i th and j th levels of the first and second factors, D_{w_1} and D_{w_2} are the directed divergences within the first and second factors and D_{B_1} and D_{B_2} are the directed divergences between the various levels of the first and second factors. Thus the total directed divergence has been decomposed into four parts, two of which correspond to directed divergence between the various levels of the first and second factors and the remaining two parts correspond to directed divergence within the various levels of the two factors.

In particular, x_{ij} and y_{ij} may represent the yields of j th variety in the i th block at two different times. The differences between x_{ij} 's and y_{ij} 's arise due to :

- (i) differences in yield between varieties
- (ii) differences in yield between blocks
- (iii) differences in yields of different varieties within the same block
- (iv) differences in yields of same variety in different blocks.

Our discussion gives the decomposition of the total directed divergence into four components, one due to each of the above four causes.!

We may also consider the special case when y_{ij} 's are all equal i. e. we may compare the present yields against the hypothesis that all blocks are equally fertile and all

varieties are equally good. Our present data do not agree fully with this hypothesis and our present discussion shows how far the differences are due to differences in varieties and how far these are due to differences in blocks.

In this case, we get

$$D_{B1} = \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \ln \frac{T_{i.}/T_{..}}{1/m} = \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \ln \frac{T_{i.}}{T_{..}} + \ln m \quad \dots(7)$$

$$D_{B2} = \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \ln \frac{T_{.j}/T_{..}}{1/n} = \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \ln \frac{T_{.j}}{T_{..}} + \frac{1}{2} \ln n \quad \dots(8)$$

$$D_{w1} = \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{x_{ij}}{T_{i.}} + \ln m \quad \dots(9)$$

$$D_{w2} = \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \sum_{i=1}^m \frac{x_{ij}}{T_{.j}} \ln \frac{x_{ij}}{T_{.j}} + \ln n \quad \dots(10)$$

$$D = \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{T_{..}} \ln \frac{x_{ij}}{T_{..}} + \ln m n \quad \dots(11)$$

so that

$$\begin{aligned} - \sum_{i=1}^m \sum_{j=1}^n \frac{x_{ij}}{T_{..}} \ln \frac{x_{ij}}{T_{..}} &= - \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \ln \frac{T_{i.}}{T_{..}} \\ &\quad - \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \ln \frac{T_{.j}}{T_{..}} - \frac{1}{2} \sum_{i=1}^m \frac{T_{i.}}{T_{..}} \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{x_{ij}}{T_{i.}} \\ &\quad - \frac{1}{2} \sum_{j=1}^n \frac{T_{.j}}{T_{..}} \sum_{i=1}^m \frac{x_{ij}}{T_{.j}} \ln \frac{x_{ij}}{T_{.j}} \quad \dots(12) \end{aligned}$$

or

$$S = \frac{1}{2} S_{B1} + \frac{1}{2} S_{B2} + \frac{1}{2} S_{w1} + \frac{1}{2} S_{w2} \quad \dots(13)$$

where S is the total entropy, S_{B_1} and S_{B_2} are the entropies between the various levels of the two factors and S_{w_1} and S_{w_2} are the weighted sums of entropies within the various levels of the two factors.

3. TESTING STATISTICAL SIGNIFICANCE OF THE TWO-FACTORS

If instead of using Kullback-Leibler¹² measure of directed divergence, we use Havrda and Charvat³ measure of degree two, and consider directed divergence from the uniform distribution, we find that D_{B_1} , D_{B_2} , D_{w_1} , D_{w_2} are the sums of squares between and within various levels, of the two factors.³

In the usual two-way analysis of variance, we pool the sum of squares within the levels of the two factors to get $D_{w_1} + D_{w_2}$ as the error sum of squares so that the number of degrees of freedom associated with the error sum of squares $D_{w_1} + D_{w_2}$ is more than the number of degrees of freedom associated with the sum of squares D_{w_1} or D_{w_2} alone. We now find the ratios of mean squares corresponding to the ratios.

$$D_{B_1}/(D_{w_1} + D_{w_2}), D_{B_2}/(D_{w_1} + D_{w_2}) \quad \dots(14)$$

and use the F test to test the significance of the two factors.

Our decomposition of directed divergence into directed divergence between and within various levels is more general than the usual decomposition of the sum of squares for the following reasons :

(i) We are using Kullback-Leibler¹² measure of directed divergence which is much more 'natural' than Havrda and Charvat³ measure of degree two

(ii) Even if we want to use Havrda and Charvat measure, we can use this measure for any degree rather than for degree two only (Kapur¹⁰).

(iii) In the usual analysis of variance, we take the second distribution implicitly as the uniform distribution. Here we can take the second distribution as any general distribution.

(iv) In particular we can consider the directed divergence (d.d) of one set of data from another set of data and decompose this d. d into d. d between and within various levels of the two factors. We can then use

$$D_{B_1}/D_{w_1}, D_{B_2}/D_{w_2}, D_{B_1}/(D_{w_1} + D_{w_2}), D_{B_2}/(D_{w_1} + D_{w_2}) \quad \dots(15)$$

as measures of the comparative influences of the two factors in the two sets of data.

To carry out a rigorous analysis of directed divergence, we have to find the exact distributions of the ratios (15) in random samples and this derivation is not going to be easy.

Even in the special case of analysis of variance, the exact distributions are available only when the parent populations are assumed to be normal. For other population distributions, the F distribution is valid only for large samples. In view of Jaynes⁴ entropy concentration theorem, the distribution of the last two ratios in¹⁵ will still be F distribution for large samples. Nevertheless knowledge of the exact distributions of these two ratios, valid for each value of n , would still be highly desirable.

4. STATISTICAL DECOMPOSITION FOR THREE OR MORE FACTORS

The above decomposition can be easily extended to the case of k factors ($k > 2$). The total directed divergence can be decomposed into $2k$ directed divergences of which k will be directed divergences between different levels of the k factors and the other k directed divergence will be weighted sums of directed divergences within the various levels of the different factors. In fact we get

$$D = \frac{1}{k} \sum_{r=1}^k D_{Br} + \frac{1}{k} \sum_{r=1}^k D_{wr} \quad \dots(16)$$

where

D_{Br} = directed divergence between the various levels of the r th factor i. e. it is the directed divergence between two probability distributions obtained from the two sets of data by dividing the total for each level by the total for all levels for that factor and

D_{wr} = weighted sums of the directed divergence within all the levels of the r th factor, then the weights being given by the total of the levels.

We similarly get

$$S = \frac{1}{k} \sum_{r=1}^k S_{Br} + \frac{1}{k} \sum_{r=1}^k S_{wr} \quad \dots(17)$$

and the measures

$$\frac{S_{Br}}{S_{Br} + S_{wr}}, \frac{D_{Br}}{D_{Br} + D_{wr}}, r = 1, 2, \dots, k. \quad \dots(18)$$

5. STATISTICAL DECOMPOSITION FOR BOTH FACTORS AND INTERACTIONS

For simplicity, we consider two factors at m and n levels and the interaction between them.

Let x_{ij_s} , y_{ij_s} be the s th observations for i th and j th levels of the two factors and let a dot suffix denote summation with respect to the variable concerned, then

$$\begin{aligned}
 \sum_{i=1}^m \sum_{j=1}^n \sum_{s=1}^N \frac{x_{ij_s}}{T_{...}} \ln \frac{x_{ij_s}/T_{...}}{y_{ij_s}/T_{...}} &= \frac{1}{3} \sum_{i=1}^m \sum_{j=1}^n \sum_{s=1}^N \frac{x_{ij_s}}{T_{...}} \\
 \ln \frac{\frac{x_{ij_s}}{T_{...}}}{\frac{y_{ij_s}}{U_{i..}}} \frac{\frac{x_{ij_s}}{T_{...}}}{\frac{y_{ij_s}}{U_{.j.}}} \frac{\frac{x_{ij_s}}{T_{...}}}{\frac{y_{ij_s}}{U_{ij.}}} \frac{\frac{T_{i..}}{T_{...}}}{\frac{U_{i..}}{U_{...}}} \frac{\frac{T_{.j.}}{T_{...}}}{\frac{U_{.j.}}{U_{...}}} \frac{\frac{T_{ij.}}{T_{...}}}{\frac{U_{ij.}}{U_{...}}} & \\
 &= \frac{1}{3} \sum_{i=1}^m \frac{T_{i..}}{T_{...}} \sum_{j=1}^n \sum_{s=1}^N \frac{x_{ij_s}}{T_{i..}} \ln \frac{x_{ij_s}/T_{i..}}{y_{ij_s}/U_{i..}} \\
 &+ \frac{1}{3} \sum_{j=1}^n \frac{T_{.j.}}{T_{...}} \sum_{i=1}^m \sum_{s=1}^N \frac{x_{ij_s}}{T_{.j.}} \ln \frac{x_{ij_s}/T_{.j.}}{y_{ij_s}/U_{.j.}} \\
 &+ \frac{1}{3} \sum_{i=1}^m \sum_{j=1}^n \frac{T_{ij.}}{T_{...}} \sum_{s=1}^N \frac{x_{ij_s}}{T_{ij.}} \ln \frac{x_{ij_s}/T_{ij.}}{y_{ij_s}/U_{ij.}} \\
 &+ \frac{1}{3} \sum_{i=1}^m \frac{T_{i..}}{T_{...}} \ln \frac{T_{i..}}{T_{...}} + \frac{1}{3} \sum_{j=1}^n \frac{T_{.j.}}{T_{...}} \ln \frac{T_{.j.}}{T_{...}} \\
 &+ \frac{1}{3} \sum_{i=1}^m \sum_{j=1}^n \frac{T_{ij.}}{T_{...}} \ln \frac{T_{ij.}/T_{...}}{U_{ij.}/U_{...}} \quad \dots(19)
 \end{aligned}$$

This is similar to break-up into three factors; the third factor in this case is the interaction between the first two factors. Thus

$$D = \frac{1}{3} (D_{w_1} + D_{B_1}) + \frac{1}{3} (D_{w_2} + D_{B_2}) + \frac{1}{3} (D_{w_{12}} + D_{B_{12}}) \quad \dots(20)$$

and

$$S = \frac{1}{3} (S_{w_1} + S_{B_1}) + \frac{1}{3} (S_{w_2} + S_{B_2}) + \frac{1}{3} (S_{w_{12}} + S_{B_{12}}). \quad \dots(21)$$

6. STATISTICAL HEIRARCHICAL DECOMPOSITION

Let x_{ijkl} y_{ijkl} be the measurements of the l th individual in the k th population in the j th region of the i th species, then

$$\begin{aligned}
 D &= \sum_{i=1}^m \sum_{j=1}^n \sum_{k=1}^p \sum_{l=1}^N \frac{x_{ijkl}}{T_{....}} \ln \frac{x_{ijkl}/T_{....}}{y_{ijkl}/U_{....}} \\
 &= \sum_{i=1}^m \sum_{j=1}^n \sum_{k=1}^p \sum_{l=1}^N \frac{T_{ijkl}}{T_{....}} \\
 &\quad \times \ln \frac{(x_{ijkl}/T_{ijkl.}) (T_{ijkl.}/T_{ij.}) (T_{ij.}/T_{i..}) (T_{i..}/T_{....})}{(y_{ijkl}/U_{ijkl.}) (U_{ijkl.}/U_{ij.}) (U_{ij.}/U_{i..}) (U_{i..}/U_{....})} \\
 &= \sum_{i=1}^m \sum_{j=1}^n \sum_{k=1}^p \frac{T_{ijkl.}}{T_{....}} \sum_{l=1}^N \frac{x_{ijkl}}{T_{ijkl.}} \ln \frac{x_{ijkl}/T_{ijkl.}}{y_{ijkl}/U_{ijkl.}} \\
 &\quad + \sum_{i=1}^m \sum_{j=1}^n \frac{T_{ij.}}{T_{....}} \sum_{k=1}^p \frac{T_{ijk.}}{T_{ij.}} \ln \frac{T_{ijk.}/T_{ij.}}{U_{ijk.}/U_{ij.}} \\
 &\quad + \sum_{i=1}^m \frac{T_{i....}}{T_{....}} \sum_{j=1}^n \frac{T_{ij.}}{T_{i..}} \ln \frac{T_{ij.}/T_{i..}}{U_{ij.}/U_{i..}} \quad \dots(22) \\
 &\quad + \sum_{i=1}^m \frac{T_{i..}}{T_{....}} \ln \frac{T_{i..}/T_{....}}{U_{i..}/U_{....}} \\
 &= D_{IV} + D_{III} + D_{II} + D_I \text{ (say)}, \quad \dots(23)
 \end{aligned}$$

where

D_I	= directed divergence between species
$D_{II} + D_{III} + D_{IV}$	= directed divergence within species
D_{II}	= directed divergence between regions (within species)
$D_{III} + D_{IV}$	= directed divergence within regions (within species)
D_{III}	= directed divergence between populations (within regions, within species)
D_{IV}	= directed divergence within populations (within regions, within species).

We can have a similar heirarchical decomposition of entropies.

7. STATISTICAL DECOMPOSITION OF HAVRDA AND CHARVAT'S ENTROPY

Let p_{ijkl} be the probability of an individual belonging to the i th religion, j th region, k th political party and l th economic status, and $p_{ijk\cdot}$ be the probability of his belonging to i th religion, j th region and k th political party and so on, then

$$\begin{aligned} \sum_i \sum_j \sum_k \sum_l p_{ijkl}^\alpha - 1 &= \left(\sum_i p_i^\alpha \dots - 1 \right) \\ &+ \sum_i p_i^\alpha \dots \left(\sum_j \left(\frac{p_{ij\cdot}}{p_{i\cdot}} \right)^\alpha - 1 \right) \\ &+ \sum_j \sum_i p_{ij\cdot}^\alpha \left(\sum_k \left(\frac{p_{ijk\cdot}}{p_{ij\cdot}} \right)^\alpha - 1 \right) \\ &+ \sum_k \sum_j \sum_i p_{ijk\cdot}^\alpha \left(\sum_l \left(\frac{p_{ijkl}}{p_{ijk\cdot}} \right)^\alpha - 1 \right). \end{aligned} \quad \dots(24)$$

By dividing by $1 - \alpha$, we get a decomposition of Havrda and Charvat's measure of entropy. Here $p_{ij\cdot}/p_{i\cdot}$ is the probability of a person of the i th religion belonging to the j th region, $p_{ijk\cdot}/p_{ij\cdot}$ is the probability of a person to belong to the k th political party when it is known that he belongs to the i th religion and j th region and $p_{ijkl}/p_{ijk\cdot}$ gives the probability of a person belonging to the l th economic status when it is known that he belongs to the i th religion, j th region and k th political party.

We can of course take religious, regions, political parties and economic strata in twenty-four different orders to get twenty-four decompositions.

8. STATISTICAL DECOMPOSITION FOR A LATIN SQUARE DESIGN

Let x_{ij} , y_{ij} be the yields from the plot in the i th row, j th column in which k th variety is sown ($i, j, k = 1, 2, \dots, n$), then

$$\begin{aligned} D &= \sum_{i=1}^n \sum_{j=1}^n \frac{x_{ij}}{T_{\cdot\cdot}} \ln \frac{x_{ij}/T_{\cdot\cdot}}{y_{ij}/U_{\cdot\cdot}} = \frac{1}{3} \sum_{i=1}^n \sum_{j=1}^n \frac{x_{ij}}{T_{\cdot\cdot}} \\ &\times \ln \frac{x_{ij}/T_{\cdot\cdot} \cdot x_{ij}/T_{\cdot j} \cdot x_{ij}/T_{k\cdot}}{y_{ij}/U_{\cdot\cdot} \cdot y_{ij}/U_{\cdot j} \cdot y_{ij}/U_{k\cdot}} \frac{T_{\cdot\cdot}/T_{\cdot\cdot} \cdot T_{ij}/T_{\cdot\cdot} \cdot T_{k\cdot}/T_{\cdot\cdot}}{U_{\cdot\cdot}/U_{\cdot\cdot} \cdot U_{ij}/U_{\cdot\cdot} \cdot U_{k\cdot}/U_{\cdot\cdot}}. \end{aligned} \quad \dots(25)$$

where T_k and U_k are the totals of the yields of k th variety in the two sets of data. Thus

$$\begin{aligned}
 D = & \frac{1}{3} \sum_{i=1}^n \frac{T_{i.}}{\bar{T}} \sum_{j=1}^n \frac{x_{ij}}{T_{i.}} \ln \frac{x_{ij}/T_{i.}}{y_{ij}/U_{i.}} \\
 & + \frac{1}{3} \sum_{j=1}^n \frac{T_{.j}}{\bar{T}} \sum_{i=1}^n \frac{x_{ij}}{T_{.j}} \ln \frac{x_{ij}/T_{.j}}{y_{ij}/U_{.j}} + \frac{1}{3} \sum_{k=1}^n \frac{T_k}{\bar{T}} \Sigma' \frac{x_{ij}}{T_k} \ln \frac{x_{ij}}{T_k} \\
 & + \frac{1}{3} \sum_{i=1}^n \frac{T_{i.}}{\bar{T}.} \ln \frac{T_{i.}}{\bar{T}.} + \frac{1}{3} \sum_{j=1}^n \frac{T_{.j}}{\bar{T}.} \ln \frac{T_{.j}}{\bar{T}.} \\
 & + \frac{1}{3} \sum_{k=1}^n \frac{T_k}{\bar{T}.} \ln \frac{T_k}{\bar{T}.} \quad \dots(26)
 \end{aligned}$$

where Σ' denotes summation over all plots which gives the k th variety. This gives the decomposition of total directed divergence into directed divergence between and within rows, columns and varieties.

9. INTERPRETATIONS FOR A CONTINGENCY TABLE

In an $m \times n$ contingency table, let O_{ij} be the observed frequency in the cell in the i th row and j th column and let a_i , b_j be the total of these frequencies in the i th row and j th column respectively, then on the basis of independence hypothesis, the expected frequency in the i - j th cell is $a_i b_j / T$, where T is the grand total of the table. The total directed divergence, between the observed frequency and the expected frequency table can then be decomposed according to (3) as follows :

$$\begin{aligned}
 \sum_{j=1}^n \sum_{i=1}^m \frac{O_{ij}}{T} \ln \frac{O_{ij}/T}{a_i b_j / T^2} = & \frac{1}{2} \sum_{i=1}^m \frac{a_i}{\bar{T}} \ln \frac{a_i/T}{a_i/\bar{T}} \\
 & + \frac{1}{2} \sum_{j=1}^n \frac{b_j}{\bar{T}} \ln \frac{b_j/T}{b_j/\bar{T}} + \frac{1}{2} \sum_{i=1}^m \frac{a_i}{\bar{T}} \sum_{j=1}^n \frac{O_{ij}}{a_i} \ln \frac{O_{ij} a_i}{b_j T} \\
 & + \frac{1}{2} \sum_{j=1}^n \frac{b_j}{\bar{T}} \sum_{i=1}^m \frac{O_{ij}}{b_j} \ln \frac{O_{ij} b_j}{a_i T}. \quad \dots(27)
 \end{aligned}$$

Since the row and column totals are the same in the two tables, there is no contribution from the directed divergences between rows and between columns

and the total directed divergences is due to k directed divergences within rows and within columns.

Similarly (12) gives the decomposition of the total entropy in the observed frequency table viz.,

$$\begin{aligned} \sum_{j=1}^n \sum_{i=1}^m \frac{O_{ij}}{T} \ln \frac{O_{ij}}{T} &= \frac{1}{2} \sum_{i=1}^m \frac{a_i}{T} \ln \frac{a_i}{T} + \frac{1}{2} \sum_{j=1}^n \frac{b_j}{T} \ln \frac{b_j}{T} \\ &+ \frac{1}{2} \sum_{i=1}^m \frac{a_i}{T} \sum_{j=1}^n \frac{O_{ij}}{a_i} \ln \frac{O_{ij}}{a_i} \\ &+ \frac{1}{2} \sum_{j=1}^n \frac{b_j}{T} \sum_{i=1}^m \frac{O_{ij}}{b_j} \ln \frac{O_{ij}}{b_j} \dots (28) \end{aligned}$$

10. DECOMPOSITION OF VECTOR OBSERVATIONS APPLICATIONS

Let X_{ij} , Y_{ij} , $T_{i.}$, $T_{.j}$, $T_{..}$, U_i , U_j , $U_{..}$ be vectors with same number of components, then the directed divergence is also a vector and each component can be decomposed as above.

Then the components of X_{ij} can be

- (i) yield of j th variety in the i th block
- (ii) fertilizer given to the j th variety in the i th block
- (iii) water given to the j th variety in the i th block
- (iv) sale price of j th variety in the i th block

All these will be highly correlated.

It may therefore be worth while to investigate the relationships between the components of the directed divergence and entropy vectors.

We may also use the normalised measure of directed divergence between two probability distribution viz.

$$\left[\sum_{i=1}^n p_i \ln \frac{p_i}{q_i} \right] / \ln \frac{1}{q_{\min}} \dots (29)$$

$$\text{where } q_{\min} = \min (q_1, q_2, \dots, q_n) \dots (30)$$

so that each component of the directed divergence vector lies between 0 and 1. Let the components of the directed divergence vector be

$$(d_1, d_2, \dots, d_n) \quad \dots(31)$$

then we form the probability distribution $p_i = d_i / \sum_{i=1}^n d_i$ and find the normalized entropy measure

$$- \sum_{i=1}^n p_i \ln p_i / \ln n. \quad \dots(32)$$

The nearer this value is to unity, the greater is the relationship between the various factors.

As another example, let $\bar{x}_{ij}, \bar{y}_{ij}$ be the mean marks of students in two schools in the i th subject in the j th class, then (3) will give the decomposition of the total directed divergence in marks of first school from second school in terms of directed divergence between classes, within classes, between subjects and within subjects.

Those are a large number of similar applications of the principle of decomposition given in the present paper to ecology, sociology, economics, anthropology genetics etc.

11. CONCLUDING REMARKS

(i) In the last section, we stated that the nearer the expression (32) is to unity, the greater will be relationship between the various factors. The question naturally arises as to how near to unity expression (32) should be. In other words, we need the distribution of $S/S_{\max}$ or of $S_{\max} - S$ where $S_{\max}$ is the maximum value of the entropy S . The answer is given by Jaynes⁴ concentration theorem which states that in $2N$ trials, $2N(S_{\max} - S)$ is distributed as chi-square with $k = n - m + 1$ degrees of freedom where m is the number of constraints other than the natural constraint. However for using the result of this theorem, we need repeated sets of observations. If we have only a single set of observations, then (32) can give only an intuitive measure of the degree of relationship between the various factors.

(ii) Equations (3), (13) can be generalised to

$$D = \lambda D_{B_1} + \lambda D_{B_2} + (1 - \lambda) D_{w_1} + (1 - \lambda) D_{w_2}; \quad 0 < \lambda < 1 \quad \dots(33)$$

$$S = \lambda S_{B_1} + \lambda S_{B_2} + (1 - \lambda) S_{w_1} + (1 - \lambda) S_{w_2}; \quad 0 < \lambda < 1. \quad \dots(34)$$

Similarly equations (20) and (21) can be generalised to

$$D = \lambda (D_{w1} + D_{B1}) + \mu (D_{w2} + D_{B2}) + \nu (D_{w12} + D_{B12}) \quad \dots(35)$$

$$S = \lambda (S_{w1} + S_{B1}) + \mu (S_{w2} + S_{B2}) + \nu (S_{w12} + S_{B12}) \quad \dots(36)$$

where

$$0 < \lambda < 1, 0 < \mu < 1, 0 < \nu < 1, \lambda + \mu + \nu = 1. \quad \dots(37)$$

Similar generalisations can be made for equations (26), (27) and (28).

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