

ON THE AVERAGE NUMBER OF CROSSINGS OF AN
ALGEBRAIC POLYNOMIAL

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(Received 31 August 1987; after revision 30 March 1988)

The present paper provides an estimate of the expected number of crossings of a polynomial of degree n with the line $y = mx$ where the coefficients are independent normally distributed and m is a constant independent of x . There are many known asymptotic estimates for the case of $m=0$. It is shown that the result is still valid even for $m \rightarrow \infty$ as long as $m = o(\sqrt{n})$.

1. INTRODUCTION

Let

$$P(x) = \sum_{i=0}^{n-1} a_i x^i \quad \dots(1.1)$$

where $a_0, a_1, a_2, \dots, a_{n-1}$ is a sequence of random variables, and $N_m(a, b) \equiv N(a, b)$ be the number of real roots of the algebraic equation $P(x) = mx$ in the interval (a, b) , where $m_n \equiv m$ is a constant independent of x , and multiple roots are counted only once. Kac⁵ found that in the case of $m = 0$, and when the coefficients of (1.1) are independent normally distributed with mean zero and variance one, the mathematical expectation of the number of real roots, $EN(-\infty, \infty)$, is asymptotic to $(2/\pi) \log n$. From later works^{3,6,8}, there is ground to believe that whatever class of distributions for the coefficients we choose it will not greatly affect the result as long as $E(a_i) = 0$ ($i = 0, 1, 2, \dots, n - 1$). Further a reduction to the expected number of real roots appears to occur in the work of Ibragimov and Maslova⁴ and Sambandham⁷ when they consider the cases of the coefficients having non zero means or being dependent. An asymptotic formula for the expected number of real roots for $P(x) = m$ is obtained by the author², which is, indeed, the number of times that $P(x)$ crosses a line parallel to the x -axis. Here we study the number of times that $P(x)$ crosses a line which is not necessarily parallel to the x -axis. We state our result in the following theorem.

Theorem—If the coefficients of (1.1) are independent, standard normal random variables, then for any constant m such that (m^2/n) tends to zero the mathematical expectation of the number of real roots of the equation $P(x) = mx$ satisfies

$$EN(-1, 1) \sim (1/\pi) \log(n/m^2) \quad \text{if } m \rightarrow \infty \text{ as } n \rightarrow \infty$$

$$EN(-1, 1) \sim (1/\pi) \log n \quad \text{if } m \text{ is bounded}$$

$$EN(-\infty, -1) = EN(1, \infty) \sim (1/2\pi) \log n.$$

From the theorem it is interesting to note that for sufficiently large n we still obtain a sizeable number of crossings even when the line tends to be perpendicular to the x -axis (i. e. $m \rightarrow \infty$).

2. A FORMULA FOR THE NUMBER OF CROSSINGS

Let $\phi(x) = (2\pi)^{-1/2} \int_{-\infty}^x \exp(-y^2/2) dy$ and $\phi(x) = \phi'(x) = (2\pi)^{-1/2} \exp(-x^2/2)$,

then by using the expected number of level crossings [(Cramer and Leadbetter¹, p. 285)] for our equation $P(x) - mx = 0$ we have

$$EN(a, b) = \int_a^b \beta^{1/2} \alpha^{-1/2} (1 - \mu^2)^{1/2} \phi(\omega/\alpha)^{1/2} [2\phi(\eta) + \eta \{2\phi(\eta) - 1\}] dx \dots (2.1)$$

where

$$\alpha(x) \equiv \alpha = \text{Var} \{P(x) - mx\}, \quad \beta(x) \equiv \beta = \text{Var} \{P'(x) - m\}$$

$$\mu = \alpha^{-1/2} \beta^{-1/2} \text{Cov} \{P(x) - mx, P'(x) - m\}, \quad \omega = E \{P(x) - mx\}$$

$$\eta = \beta^{-1/2} (1 - \mu^2)^{-1/2} \{v - \beta^{1/2} \mu \omega \alpha^{-1/2}\} \text{ and } v = E \{P'(x) - m\}.$$

Since the coefficients of $P(x)$ are independent random variables with mean zero and variance one we can easily find that

$$\omega = -mx, \quad \alpha = \text{Var} \{P(x)\} = \sum_{i=1}^{n-1} x^{2i}$$

$$\beta = \text{Var} \{P'(x)\} = \sum_{i=1}^{n-1} i^2 x^{2i-2}, \quad v = -m$$

$$\mu = \alpha^{-1/2} \beta^{-1/2} E \{P(x) \cdot P'(x)\} = \alpha^{-1/2} \beta^{-1/2} \sum_{i=1}^{n-1} i x^{2i-1} = \alpha^{-1/2} \beta^{-1/2} \gamma$$

(say)

and

$$\eta = m \alpha^{-1/2} (\alpha\beta - \gamma^2)^{-1/2} (\gamma x - \alpha).$$

Hence from (2.1) and since $\phi(x) = \frac{1}{2} + (2\pi)^{-1/2} \text{erf}(x)$ we have the extension of Kac-Rice formula⁹.

$$EN(a, b) = \int_a^b (\Delta^{1/2}/\pi\alpha) \exp \{(-am^2 + 2m^2 \gamma x - \beta m^2 x^2)/2\Delta\}$$

(equation continued on p. 3)

$$\begin{aligned}
 & + \{ | m (\gamma x - \alpha) | / \pi \alpha^{3/2} \} \exp (-m^2 x^2 / 2\alpha) \operatorname{erf} \\
 & \{ | m (\gamma x - \alpha) | / \alpha^{1/2} \Delta^{1/2} \} dx \\
 = & \int_a^b I(x) dx = \int_a^b I_1(x) dx + \int_a^b I_2(x) dx \text{ (say)} \quad \dots(2.2)
 \end{aligned}$$

where

$$\Delta(x) \equiv \Delta = \alpha \beta - \gamma^2.$$

3. PROOF OF THE THEOREM

First we consider the case when $m \rightarrow \infty$ as $n \rightarrow \infty$. For finding an upper limit of $EN(0, 1)$, let $0 \leq a \leq 1$ and divide $(0, 1)$ into two sub-intervals $(0, 1 - 1/n^a)$ and $(1 - 1/n^a, 1)$. For $0 \leq x \leq 1 - 1/n^a$ and all sufficiently large n we have

$$\begin{aligned}
 \gamma & = \{(n - 1) x^{2n+1} - n x^{2n-1} + x\} (1 - x)^{-2} \\
 & = x (1 - x^{2n}) (1 - x^2)^{-2} + O \{n^{1+a} \exp(-2n^{1-a})\} \quad \dots(3.1)
 \end{aligned}$$

and

$$\beta = (1 + x^2) (1 - x^{2n}) (1 - x^2)^{-3} + O \{n^{2+a} \exp(-2n^{1-a})\}. \quad \dots(3.2)$$

From (3.1) and (3.2) we have

$$\Delta = (1 - x^{2n}) (1 - x^2)^{-4} + O \{n^{2+2a} \exp(-2n^{1-a})\}. \quad \dots(3.3)$$

Now we choose $a = 1 - \log \log (n^{10}) / \log n$. Then since

$$\exp(-2n^{1-a}) \exp\{-2 \log(n)^{10}\} = n^{-20}$$

for n sufficiently large, all the terms inside the $O\{\}$ will tend to zero and also $a \rightarrow 1$ which is necessary for obtaining our further result. Hence from (2.2), (3.1), and (3.3) for all sufficiently large n we have

$$\begin{aligned}
 \int_0^{1-1/n^a} I_1(x) dx & = \int_0^{1-1/n^a} (\Delta^{1/2} \pi \alpha) \exp \{[-m^2 (1 - x^2) / (1 - x^{2n})] \{(1 - x^2)^2 / 2 \\
 & - x^2 (1 - x^2) + x^2 (1 + x^2) / 2\} \{1 + O(n^{2+a} \\
 & \times \exp(-2n^{1-a}))\}\} dx. \quad \dots(3.4)
 \end{aligned}$$

Now we note that for all x

$$\begin{aligned}
 (1 - x^2)^2 / 2 - x^2 (1 - x^2) + x^2 (1 + x^2) / 2 & = (1 - 3x^2 + 4x^4) / 2 \\
 & \geq 7/32 > 1/5
 \end{aligned}$$

and since from Kac⁵ we have $\Delta^{1/2} / \alpha < (1 - x^2)^{-1}$, (3.4) gives

$$\int_0^{1-1/n^a} I_1(x) dx < (1/\pi) \int_0^{1-1/n^a} (1 - x^2)^{-1} \exp\{-m^2 (1 - x^2) / 5 (1 - x^{2n})\} dx$$

(equation continued on p. 4)

$$\begin{aligned}
 &< (1/\pi) \int_0^{1-1/n^a} (1-x^2)^{-1} \exp\{-m^2(1-x^2)/5\} dx. \\
 &\dots(3.5)
 \end{aligned}$$

Let $\lambda = m^2/5$, then from (3.5) and since $\exp\{-\lambda(1-x^2)\} < \{1 - \lambda(1-x^2)^{-1}\}$ we have

$$\begin{aligned}
 \int_0^{1-1/n^a} I_1(x) dx &< (1/\pi) \int_0^{1-1/n^a} (1-x^2)^{-1} \{1 + \lambda(1-x^2)^{-1}\} dx \\
 &= (1/\pi) \int_0^{1-1/n^a} [(1-x^2)^{-1} - \lambda\{1 + \lambda(1-x^2)^{-1}\}] dx \\
 &= (1/2\pi) \log \frac{2 - 1/n^a}{1/n^a} - (1/2\pi) (1 - (1/\lambda)^{-1/2}) \\
 &\quad \log \frac{\sqrt{1 + (1/\lambda)} + 1 - (1/n^a)}{\sqrt{1 + (1/\lambda)} - 1 + (1/n^a)}. \dots(3.6)
 \end{aligned}$$

Now since $(1 + 1/\lambda)^{1/2} = 1 + 1/2\lambda + O(1/\lambda^2)$ for all $|\lambda| < 1$ and also by noting that $a \rightarrow 1$, $\lambda \rightarrow \infty$ and $\lambda/n \rightarrow 0$ as $n \rightarrow \infty$, from (3.6) for all sufficiently large n we have

$$\begin{aligned}
 \int_0^{1-1/n^a} I_1(x) dx &= (1/2\pi) (\log 2 + \log n) \\
 &\quad - (1/2\pi) (1 + 1/\lambda)^{-1/2} \log \frac{2 + 1/2\lambda + O(1/\lambda^2) - 1/n}{1/2\lambda + O(1/\lambda^2) + 1/n} \\
 &< (1/2\pi) (\log 2 + \log n) - (1/2\pi) (1 + 1/\lambda)^{-1/2} \\
 &\quad \{\log \lambda + \log(4 - 1/n)\} < (1/2\pi) \log n/m^2 + 1. \\
 &\dots(3.7)
 \end{aligned}$$

Also for $1 - 1/n^a \leq x \leq 1$, since from Ibragimov and Maslova⁴ $\Delta^{1/2}/\alpha < (2n-1)^{1/2} (1-x)^{-1/2}$, we have

$$\begin{aligned}
 \int_{1-1/n^a}^1 I_1(x) dx &< \int_{1-1/n^a}^1 (2n-1)^{1/2} (1-x)^{-1/2} dx \\
 &= 2(2n^{1-a} - n^{-a})^{1/2} = O(\log n)^{1/2}. \dots(3.8)
 \end{aligned}$$

Now we find an upper limit for $\int_0^1 I_2(x) dx$. We have

$$mx\gamma\alpha^{-3/2} = mx(1-x^2)^{-1/2} \{(n-1)x^{2n+1} - nx^{2n-1} + x\} (1-x^{2n})^{-3/2} \dots(3.9)$$

which gives

$$|m(\gamma x - \alpha) \alpha^{-3/2}| = |m| (1-x^2)^{-1/2} (1-x^{2n})^{-3/2} |(2x^2 - 1)(1-x^{2n}) - n x^{2n} (1-x^2)|. \quad \dots(3.10)$$

Now for $0 \leq x \leq \sqrt{3/2}$ given any ϵ positive exists an integer n_0 such that for all $n \geq n_0$

$$n x^{2n} (1-x^2) < n (3/4)^n < \epsilon.$$

We note that $(2x^2 - 1)(1-x^{2n}) < (2x^2 - 1) < \epsilon$ only for a small interval of $\frac{1}{2}(1-\epsilon) < x^2 < \frac{1}{2}(1+\epsilon)$. Hence from (3.10) for all sufficiently large n we have

$$\int_{\sqrt{(1-\epsilon)/2}}^{\sqrt{(1+\epsilon)/2}} I_2(x) dx < |m| (\sqrt{2/\pi}) \{1 - (3/4)^n\} \{\epsilon^2 + O(\epsilon^3)\}. \quad \dots(3.11)$$

Using the fact that ϵ can be made arbitrarily small we can see that the above integral tends to zero as $n \rightarrow \infty$ (for example choosing $\epsilon = 1/m$). On the other hand, let $\int_0^{1-1/n}$

$\dots dx$ indicate the integral over $0 < x \leq 1 - 1/n$ excluding $\sqrt{(1-\epsilon/2)} \leq x \leq \sqrt{(1+\epsilon)/2}$, then from (3.10) and since for $0 \leq x \leq 1 - 1/n$ and n sufficiently large $x^{2n} < (1 - 1/n)^{2n} \rightarrow e^{-2}$, we have

$$\begin{aligned} \int_0^{1-1/n} I_2(x) dx &\leq |m| (\sqrt{2/\pi}) (1 - e^{-2})^{-1/2} \int_0^{1-1/n} (1 - 2x^2) (1 - x^2)^{-1/2} \\ &\quad \exp\{-m^2 x^2 (1 - x^2)\} \leq |m| (\sqrt{2/\pi}) (1 - e^{-2})^{-1/2} \int_0^{1-1/n} \\ &\quad (1 - 2x^2) (1 - x^2)^{-1/2} \exp\{-m^2 x^2 (1 - x^2)\} dx \\ &< (2\pi)^{-1/2} (1 - e^{-2})^{-1/2} \int_0^k u^{-1/2} e^{-u} du \end{aligned}$$

where $u = m^2 x^2 (1 - x^2)$ and $k = m^2 (1 - 1/n)^2 (2/n - 1/n^2)$. Now by integrating by parts and since $u^{1/2} e^{-u} < 1$ for all sufficiently large n , we have

$$\begin{aligned} \int_0^{1-1/n} I_2(x) dx &< (2\pi)^{-1/2} (1 - e^{-2})^{-1/2} \{[2(m^2/n)(2 - 1/n)(1 - 1/n)^2]^{1/2} \\ &\quad \exp\{-(m^2/n)(2 - 1/n)(1 - 1/n)^2\} + 2(m^2/n)(1 - 1/n)^2 \\ &\quad (2 - 1/n)\} \quad \dots (3.12) \end{aligned}$$

which also tends to zero as $n \rightarrow \infty$. In order to estimate $\int_{1+1/n}^1 I_2(x) dx$ we note that

always $\gamma > \alpha/x$, then

$$|\gamma x/\alpha^{3/2} - 1/\alpha| < 2 |\gamma x/\alpha^{3/2}|. \quad \dots(3.13)$$

Since

$$\gamma + \sum_{i=1}^{n-1} i x^{2n-1} < (n/x) \sum_{i=1}^{n-1} x^{2i} = (n/x) \alpha$$

for $1 \leq x \leq 1 - 1/n$ we can obtain

$$\gamma x/\alpha^{3/2} \leq n \alpha^{-1/2} \leq \sqrt{n} (1 - 1/n)^{-(n-1)}.$$

This and (3.13) gives

$$\begin{aligned} \int_{1-1/n}^1 I_2(x) dx &< (2 |m| / \pi) \int_{1-1/n}^1 (\gamma x/\alpha^{3/2}) dx \\ &< (2 |m| / \pi \sqrt{n}) (1 - 1/n)^{-(n-1)} \end{aligned} \quad \dots(3.14)$$

which also tends to zero as $n \rightarrow \infty$. From (3.7), (3.8), (3.11), (3.12) and (3.14) we have

$$EN(0, 1) < (1/2\pi) \log(n/m^2) + O\{\log n\}^{1/2}. \quad \dots(3.15)$$

In order to obtain a lower estimate of $EN(0, 1)$ without loss of the generality we can assume $2m^2 > 1$. Then from (3.4) and since for all $0 \leq x \leq 1$

$$(1 - x^2)^2/2 - x^2(1 - x^2) + x^2(1 + x^2)/2 \leq 1$$

we have

$$\begin{aligned} \int_0^{1-1/n^a} I_1(x) dx &> (1/\pi) \int_0^{1-1/n^a} (\Delta^{1/2}/\alpha) \exp\{-m^2(1 - x^2)/(1 - x^{2n})\} dx \\ &\geq (1/2\pi) \int_0^{1-1/n^a} (1 - x)^{-1} \exp\{-2m^2(1 - x)\} dx \\ &= (1/2\pi) \int_{2m^2/n^a}^{2m^2} u^{-1} e^{-u} du = (1/2\pi) \left\{ \int_{2m^2/n^a}^{2m^2} u^{-1} du \right. \\ &\quad \left. - \int_{2m^2/n^a}^1 u^{-1} (1 - e^{-u}) du - \int_1^{2m^2} u^{-1} du + \int_1^{2m^2} u^{-1} e^{-u} du \right\} \end{aligned} \quad \dots(3.16)$$

where $u = 2m^2(1 - x)$. Now by noticing that $u^{-1}(1 - e^{-u}) < 1$ and $\int_1^{2m^2} u^{-1} e^{-u} du$ is always positive, from (3.16) for all sufficiently large n we have

$$\begin{aligned} \int_0^{1-1/n^a} I_1(x) dx &> (1/2\pi) \{a \log n - \log m^2 - 1 - \log 2\} \\ &= (1/2\pi) \log(n/m^2) + O(\log \log n). \end{aligned} \quad \dots(3.17)$$

Finally from (3.15) and (3.17) we obtain the asymptotic formula

$$EN(0, 1) \sim \log(n/m^2).$$

On the other hand for m bounded from (3.5) and (3.8) we have

$$\begin{aligned} \int_0^1 I_1(x) dx &< (1/\pi) \int_0^{1-1/n^a} (1-x^2)^{-1} dx + (1/\pi) \int_{1-1/n^a}^1 (2n-1)^{1/2} (1-x)^{-1/2} \\ &\times dx < (1/2\pi) \log n + O\{\log n\}^{1/2}. \end{aligned}$$

From this and (3.11), (3.12) and (3.14) we have

$$EN(0, 1) < (1/2\pi) \log n + O\{\log n\}^{1/2}. \quad \dots(3.18)$$

Also in order to obtain a lower estimate, from (3.16) for all sufficiently large n we have

$$\begin{aligned} \int_0^{1-1/n^a} I_1(x) dx &> (a/2\pi) \log n - 2m^2 \\ &= (1/2\pi) \log n + O\{\log \log n\}. \end{aligned} \quad \dots(3.19)$$

Hence from (3.18) and (3.19) for m bounded we have

$$EN(0, 1) \sim (1/2\pi) \log n.$$

Now we shall find the asymptotic relation for $EN(1, \infty)$. Let $y = 1/x$ then we have

$$\int_1^\infty I(x) dx = \int_0^1 I(1/y) y^{-2} dy. \quad \dots(3.20)$$

Since from Farahmand² (p. 707) we have $\Delta(1/y) = y^{-2(2n-4)} \Delta(y)$ and $\alpha(1/y) = y^{-(2n-2)} \alpha(y)$, from (2.2) and (3.20) we have

$$\begin{aligned} \int_1^\infty I_1(x) dx &< (1/\pi) \int_0^1 y^2 \{\Delta(1/y)\}^{1/2} / \alpha(1/y) dy \\ &= (1/\pi) \int_0^1 \{\Delta(y)\}^{1/2} / \alpha(y) dy \\ &< (1/\pi) \int_0^{1-1/n} (1-y^2)^{-1} dy + (1/\pi) \int_{1-1/n}^1 (2n-1)^{1/2} (1-y)^{-1/2} dy \\ &< (1/2\pi) \log n + 3.5. \end{aligned} \quad \dots(3.21)$$

Now we find an upper estimate for $I_2(x)$. We have

$$\begin{aligned} \{(m/y) \gamma(1/y) - m(\alpha(1/y))\} \{\alpha(1/y)\}^{-2/2} \\ = m y^{n-1} (1-y^2)^{-1/2} (1-y^n)^{-3/2} \{n(1-y^2) - (2-y^2)(1-y^{2n})\}. \end{aligned} \quad \dots(3.22)$$

Let a be the same constant as before. From (2.2), (3.13) and (3.22) and for all sufficiently large n we have

$$\begin{aligned} \int_0^{1-1/n^a} y^{-2} I_2(1/y) dy &\leq (2 |m| n/\pi) \int_0^{1-1/n^a} y^{n-3} (1-y^2)^{1/2} (1-y^{2n})^{-3/2} dy \\ &\leq (2 |m| n/\pi) \exp(-n^{1-a}) \{1 - \exp(2n^{1-a})\}/(n-2) \quad \dots (3.23) \end{aligned}$$

which, since $\exp(-n^{1-a}) = n^{-1^0}$, tends to zero as $n \rightarrow \infty$. On the other hand for $1 - (1/n^a) \leq y \leq 1$

$$\begin{aligned} x^\gamma(x) \{\alpha(x)\}^{-3/2} &< n \{\alpha(x)\}^{-1/2} = n y^{n-1} (1-y^2)^{1/2} (1-y^{2n})^{-1/2} \\ &< n y^{n-1} \{2n^{-a} - n^{-2a}\}^{1/2} \{1 - (1 - (1/n^a)^{2n})\}^{-1/2}. \end{aligned}$$

Hence from this, (2.2) and (3.13) we have

$$\begin{aligned} \int_{1-1/n^a}^1 y^{-2} I_2(1/y) dy &\leq (2 |m| n\sqrt{2/\pi}) n^{-a/2} \{1 - (1/n^a)^{2n}\}^{-1/2}/n - 2 \\ &= O\{\log n\}^{1/2}. \quad \dots(3.24) \end{aligned}$$

Since $n^{-a/2} = (n^{-1} \log n^{10})^{1/2}$. Hence from (3.21), (3.23) and (3.24) we obtain

$$EN(1, \infty) < (1/2\pi) \log n + O\{\log n\}^{1/2}. \quad \dots(3.25)$$

To get the lower estimate for $EN(1, \infty)$, for $0 \leq y \leq 1 - 1/n^a$ we have

$$\alpha = y^{-(2n-2)} (1 - y^{2n})/(1 - y^2), \quad \gamma = y^{-(2n-3)} \{1 - (1 - y^{2n})/(1 - y^2)\}/(1 - y^2),$$

$$\begin{aligned} \Delta &= y^{-(4n-8)} \{(1 - y^{2n})^2 (1 - y^2)^{-4} - n^2 y^{2n-2} (1 - y^2)^{-2}\} \\ &= y^{-(4n-8)} (1 - y^{2n})^2 (1 - y^2)^{-4} + O\{n^{2+2a} \exp(-2n^{1-a})\} \end{aligned}$$

and

$$\begin{aligned} \beta/\Delta &= y^{2n-4} (1 - y^2) (1 - y^{2n})^{-1} \{1 + y^2 + n^2 (1 - y^2)^2 (1 - y^{2n})^{-1} \\ &\quad - 2n (1 - y^2) (1 - y^{2n})\} + O\{n^{2+2a} \exp(-2n^{1-a})\}, \end{aligned}$$

Hence, from these, and simple algebra, we find

$$\begin{aligned} (-\alpha + 2\gamma/y - \beta/y^2)/\Delta &= y^{2n-6} (1 - y^2) (1 - y^{2n})^{-1} [(1 - y^2)^2 \\ &\quad - 2 (1 - y^2)^2 (1 - y^{2n})^{-1} \{1 - (1 - y^{2n})/(1 - y^2)\} + 1 + y^2 + n^2 \\ &\quad (1 - y^2)^2/(1 - y^{2n}) - 2n (1 - y^2)/(1 - y^{2n})] \\ &\quad + O\{n^{2+2a} \exp(-2n^{1-a})\} < n^2 \exp(2n^{1-a}) = n^{-18} \quad \dots(3.26) \end{aligned}$$

for all sufficiently large n . Now, since from (3.26), $\exp\{m^2(-\alpha + 2\gamma/y - \beta/y^2)/2\Delta\}$ tends to one as $n \rightarrow \infty$, from (3.20), (3.26) and for all sufficiently large n

$$\begin{aligned}
 EN(1, \infty) &> \int_1^{\infty} I_1(x) dx > (1/\pi)^{1-1/n^a} \int_0^1 (1-y^2)^{-1} dy \\
 &= (a/2\pi) \log n - (1/2\pi) \log(2 - 1/n^a) = (1/2\pi) \log n \\
 &\quad + O(\log \log n). \qquad \dots(3.27)
 \end{aligned}$$

Finally from (3.25) and (3.27) we have asymptotic formula

$$EN(1, \infty) \sim (1/2\pi) \log n.$$

Since a_j and $-a_j$ ($j = 0, 1, \dots, n-1$) both have the standard normal distribution

$EN(-1, 0) = EN(0, 1)$ and $EN(-\infty, -1) = EN(1, \infty)$, we have proof of the theorem.

If we add this result to asymptotic formula in Farahmand² for k constant such that (k^2/n) tends to zero we can obtain the result for the case in which $P(x) = mx + k$. For this case

$$\omega = -mx - k \text{ and } \eta = \{\gamma(mx + k) - \alpha m\} / \alpha^{1/2} \Delta^{1/2}$$

while α, β, γ and μ will remain the same. The Kac-Rice formula will be

$$\begin{aligned}
 EN(a, b) &= \int_a^b (\Delta^{1/2}/\pi\alpha) [\exp\{-\alpha m^2 + 2m\gamma(mx + k) - \beta(mx + k)^2\}/2\Delta] \\
 &\quad + \{[-\alpha m + \gamma(mx + k)]/\pi\alpha^{3/2}\} \exp\{-(mx + k)^2/2\alpha\} \\
 &\quad \times \operatorname{erf}\{[-\alpha m + \gamma(mx + k)]/\alpha^{1/2} \Delta^{1/2}\} dx
 \end{aligned}$$

and it is interesting to know for this case

$$EN(-1, 1) \sim (1/\pi) \log(n/(m^2 + k^2)) \text{ if either } m \text{ or } k \text{ or both tend to infinity as } n \rightarrow \infty$$

$$EN(-1, 1) \sim (1/\pi) \log n \quad \text{if } m \text{ and } k \text{ are bounded}$$

$$EN(-\infty, 1) = EN(1, \infty) \sim (1/2\pi) \log n.$$

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