

# AN EXTENSION OF THE FIXED POINT PROBABILITY VECTOR FOR THE SUCCESSIVE COMPOSITION OF TRANSITION MATRICES

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(Received 9 March 1977; after revision 30 May 1977)

In this paper we have tried to establish a fixed point probability vector for the successive composition of a finite set of transition matrices—an extension of Heenan's result (1970). An example is given to show that this condition fails in the case of a finite product of such like matrices

The well-known Banach contraction principle has been used in the field of fixed points in various forms. Browder (1966) applied this principle for various types of operators, linear or non-linear, acting on various spaces to evaluate the fixed point property. Belluce and Kirk (1969) used it for non-expansive mappings in Banach space. Heenan (1970) devised a method to determine the fixed point in a transition matrix. He showed that if  $A$  is the transpose of a probability matrix  $W$ , then  $\exists w \in W$ , such that  $A(w) = w$ . By a transpose of a probability matrix  $W$ , we mean a matrix in which the sum of each column is unity.

The object of the present study is to extend the above result for the set of non-singular transition matrices which reduces Heenan's result (1970) as a corollary. Examples are cited to illustrate the above fact and its failure. Throughout the paper,  $W$  stands for the set of matrices under composition and  $A_i$ 's denote the transpose of the corresponding transition matrices.

§1. The following definition of Markov system will be required to explain the concept of transition matrix.

*Definition 1.1* — Let us consider a probabilistic system which can assume any one of the finite number of states, say  $N$ . These states are all distinct and identifiable and thus can be ordered by an index  $k$ , where  $k = 1, 2, \dots, N$ . At the start of a certain prescribed interval, the system may or may not change from its present state to one of the other  $N - 1$  possible states. The start of these intervals at which the transitions take place need not be identifiable with equal time intervals. There exists a certain class of stochastic processes for which the transition probability

$p_{ij}$  for the system to change from state  $i$  to state  $j$  at the beginning of one of these intervals is constant and depends only upon  $i$  and  $j$ . In other words, the transition probability is independent of the past history of the system, equivalently, any past transitions. Such a stochastic model is called a simple Markov system.

We now define the transition matrix.

*Definition 1.2* — The transition matrix, which is also known as a stochastic matrix, is the mathematical representation of Markov system where each element in the matrix  $p_{ij}$  is the probability that a system now in state  $i$  will transit to state  $j$  at the beginning of the next interval. Thus, the matrix becomes

$$\begin{bmatrix} p_{11} & p_{12} & \dots & p_{1N} \\ p_{21} & p_{22} & \dots & p_{2N} \\ \dots & \dots & \dots & \dots \\ p_{N1} & p_{N2} & \dots & p_{NN} \end{bmatrix} \quad \dots(1.1)$$

where  $0 \leq p_{ij} \leq 1; \quad i, j = 1, 2, \dots, N$

and  $\sum_{i=1}^N p_{ii} = 1; \quad i = 1, 2, \dots, N.$

The terms  $p_{ii}$  denote the possibilities for the system to remain at the same state from one interval to another.

If we define the probability that a Markov system is in state  $i$  during the  $n$ th period as  $P_i(n)$ , then

$$P_i(n + 1) = \sum_{i=1}^N p_{ij} P_i(n), \quad j = 1, 2, \dots, N. \quad \dots(1.2)$$

There are two properties that a simple Markov system possesses in most of the cases :

- (1) The limiting value of  $P_i(n)$  ( $i = 1, 2, \dots, N$ ), as  $n \rightarrow \infty$  is unique. These values are called steady-state probabilities of the Markov process.
- (2) The limiting value of a property is independent of the state from which the system initially starts at  $n = 0$ .

§2. In this section, we proceed to prove the following theorem.

*Theorem 2.1* — Let  $A_1, A_2, A_3, \dots, A_n$  be the finite set of non-singular matrices and  $W$  be the set of corresponding probability matrices, then  $\exists w \in W$ , such that

$$A_1 \circ A_2 \circ \dots \circ A_n(w) = w. \quad \dots(2.1)$$

PROOF : To prove (2.1), we shall employ the technique of mathematical induction.

The theorem is obviously true, when  $n = 1$ .

Now supposing that theorem (2.1) is true for  $n$ , we shall show that it is also true for  $(n + 1)$ . We have

$$A_1 \circ A_2 \circ A_3 \circ \dots \circ A_{n+1}(w) = A_1 \circ (A_2 \circ A_3 \circ \dots \circ A_{n+1})(w). \quad \dots(2.2)$$

Using the property that every non-singular matrix  $A$  can be reduced to the form  $PQ^{-1}$ , where  $P$  is symmetrical and  $Q$  is diagonal, each element of  $A_i$ 's can be written in the following ratio form :

$$a_{ij}^1 = \frac{p_{ij}^1}{q_i^1}, a_{ij}^2 = \frac{p_{ij}^2}{q_i^2}, \dots, a_{ij}^n = \frac{p_{ij}^n}{q_i^n} \quad \dots(2.3)$$

where  $q_i^k$ 's  $\in Q$ ,  $k = 1, 2, \dots, n$ ,

$$i, j = 1, 2, \dots, N.$$

At the steady state, the components of vector  $w$  are given by

$$w_i^k = \frac{q_i^k}{\sum_{j=0}^{N-1} q_j^k}, k = 1, 2, \dots, n; i = 0, 1, 2, \dots, N - 1. \quad \dots(2.4)$$

In view of relations (2.3) and (2.4), we get

$$\begin{aligned} A_1 \circ (A_2 \circ A_3 \circ \dots \circ A_{n+1})(w) &= \\ &= \sum_{i,j=0}^{N-1} \frac{p_{ij}^1}{q_i^1} w_i \circ \sum_{i,j=0}^{N-1} \frac{p_{ij}^k}{q_i^k} w_i^k \quad (k = 2, 3, \dots, (n + 1)) \\ &= \sum_{i,j=0}^{N-1} \frac{p_{ij}^{k+1}}{q_i^{k+1}} w_i^{k+1} \quad (k = 1, 2, \dots, n) \\ &= w. \end{aligned}$$

(Here the  $q_i$ 's common to each column will be identical, because all  $A_i$ 's are non-singular matrices of the set of corresponding probability matrices  $W$ ).

This completes the proof of the theorem.

The components of the probability vector  $w$  in terms of  $p_{ij}^k$  and  $q_j^k$  can be obtained from the following relations.

$$\sum_{i=0}^{N-1} \frac{p_{ij}^k}{q_i^k} w_i^k = w_j^k$$

where  $k = 1, 2, \dots, n; j = 0, 1, 2, \dots, N - 1.$  ... (2.5)

Further, if

$$w_j^k = \frac{q_j^k}{\sum_{i=0}^{N-1} q_i^k} \quad \begin{matrix} i, j = 0, 1, 2, \dots, N - 1 \\ k = 1, 2, \dots, n, \end{matrix}$$

then (2.5) becomes

$$\sum_{i=0}^{N-1} p_{ij}^k = q_j^k \quad (k = 1, 2, \dots, n) \quad \dots(2.6)$$

which is valid if the matrices  $p_i$ 's are symmetric.

*Corollary* — If  $A_1 = A_2 = A_3 = \dots = A_n$  be the finite set of non-singular matrices, and  $W$  be the set of corresponding ( $w_1 = w_2 = \dots = w_n = W$ ) probability matrices, then  $\exists w \in W$ , such that

$$A^n w = w. \quad \dots(2.7)$$

**Illustrations**

*Example 1* — To prove  $A_1 \circ A_2(w) = w$ , where

$$A_1 = \begin{bmatrix} 0.7 & 0.2 & 0.1333 \\ 0.1 & 0.6 & 0.0667 \\ 0.2 & 0.2 & 0.8 \end{bmatrix} \quad \text{and} \quad A_2 = \begin{bmatrix} 0.67 & 0.22 & 0.147 \\ 0.11 & 0.56 & 0.073 \\ 0.22 & 0.22 & 0.78 \end{bmatrix}$$

They can be reduced to the form  $p^i Q^{i-1}$ , where  $i = 1, 2, \dots$

$$A_1 = \begin{bmatrix} 7 & 1 & 2 \\ 1 & 3 & 1 \\ 2 & 1 & 12 \end{bmatrix} \begin{bmatrix} 1/10 & & \\ & 1/5 & \\ & & 1/15 \end{bmatrix}$$

and  $A_2 = \begin{bmatrix} 67 & 11 & 22 \\ 11 & 28 & 11 \\ 22 & 11 & 117 \end{bmatrix} \begin{bmatrix} 1/100 & & \\ & 1/50 & \\ & & 1/150 \end{bmatrix}$

Here,  $q_0^1 = 10$ ,  $q_1^1 = 5$ ,  $q_2^1 = 15$ ,  $\therefore \Sigma q^1 = 30$  and

$$w_0^1 = 10/30 = 1/3, w_1^1 = 5/30 = 1/6, w_2^1 = 15/30 = 1/2. \text{ And}$$

$$q_0^2 = 100, q_1^2 = 50, q_2^2 = 150, \therefore \Sigma q^2 = 300,$$

$$w_0^2 = 100/300 = 1/3, w_1^2 = 50/300 = 1/6, w_2^2 = 150/300 = 1/2.$$

Hence,

$$A_1 \circ A_2(w) = w.$$

The physical interpretation of  $A_1 \circ A_2$  is as follows. Out of the total 30 and 300 events respectively, the process in the state 0, state 1 and state 2, is 10, 5 and 15 times for  $A_1$  and 100, 50 and 150 times for  $A_2$  respectively.

The theorem is not valid if we replace the composition of the matrices by the finite product of matrices, as is obvious from the following example.

*Example 2* — To prove  $A_1 \times A_2(w) \neq w$ , consider  $A_1, A_2$  as given in Example 1. We have

$$A_1 \times A_2 = \begin{bmatrix} 0.7 & 0.2 & 0.1333 \\ 0.1 & 0.6 & 0.0667 \\ 0.2 & 0.2 & 0.8 \end{bmatrix} \times \begin{bmatrix} 0.67 & 0.22 & 0.147 \\ 0.11 & 0.56 & 0.073 \\ 0.22 & 0.22 & 0.78 \end{bmatrix}$$

which on simplification reduces to

$$\begin{bmatrix} 0.520326 & 0.295326 & 0.221474 \\ 0.147647 & 0.372674 & 0.110526 \\ 0.332 & 0.332 & 0.9326 \end{bmatrix}$$

Here,  $P$  is no longer symmetric. Thus,

$$A_1 \times A_2(w) \neq w.$$

#### REFERENCES

- Browder, F. E. (1966). Fixed point theorem for non-linear semi-contractive mapping in Banach spaces. *Archs ration. Mech. Analysis*, **21**, 259-69.
- Belluce, L. P., and Kirk, W. (1969). Fixed point theorem for certain classes of non-expansive mapping. *Proc. Am. math. Soc.*, **2**, 141-46.
- Heenan, N. I. (1970). Determination of the fixed point probability vector of a transition matrix. *Proc. IEEE*, 799.
- Wadsworth George, P., and Bryan Joseph, G. (1974). Applications of Probability Variable. McGraw-Hill Book Co., New York.